SCIENTIFIC REPORTS
nature research

OPEN

# Genome-wide association mapping and accuracy of predictions for amoebic gill disease in Atlantic salmon (*Salmo salar*)

**Muhammad L. Aslam[1*], Solomon A. Boison[1,2], Marie Lillehammer[1], Ashie Norris[2] & Bjarne Gjerde[1]**

**Amoebic gill disease (AGD) is a parasitic disease caused by the amoeba *Paramoeba perurans*, which colonizes the gill tissues and causes distress for the host. AGD can cause high morbidity and mortalities in salmonid and non-salmonid fish species. To understand the genetic basis of AGD and improve health status of farmed A. salmon, a population of ~ 6,100 individuals belonging to 150 full-sib families was monitored for development of AGD in the sea of Ireland. The population was followed for two rounds of AGD infections, and fish were gill scored to identify severity of disease in first (N = 3,663) and the second (N = 3,511) infection with freshwater treatment after the first gill-scoring. A subset of this gill-scored population (N = 1,141) from 119 full-sib families were genotyped with 57,184 SNPs using custom-made Affymetrix SNP-chip. GWAS analyses were performed which resulted in five significantly associated SNP variants distributed over chromosome 1, 2 and 5. Three candidate genes; *c4*, *tnxb* and *slc44a4* were found within QTL region of chromosome 2. The *tnxb* and *c4* genes are known to be a part of innate immune system, and may play a role in resistance to AGD. The gain in prediction accuracy obtained by involving genomic information was 9–17% higher than using traditional pedigree information.**

Amoebic gill disease (AGD) is caused by a parasite *Paramoeba perurans*, which colonizes gill tissue[1,2] and ultimately causes inappetence, respiratory distress and cardiovascular compromise[3,4]. Attachment of amoebae to the gill initiates a localized host cellular response, including hyperplasia and hypertrophy of the gill epithelium and lamellar fusion[5]. This pathological condition can cause high production losses in multiple salmonid and non-salmonid fish species[3,4]. Although AGD can occur year-round, it is most prevalent in warmer water and high salinity with increased frequency and severity[6]. The existence and severity of disease at farmed facilities is evaluated by random sampling of fish (~10–15) and investigation and/or scoring of gills for pathological conditions caused by amoeba as described by Taylor *et al.*[7]. The detection for the specific type of pathogen/amoeba and the pathogen load (measure of disease severity) can also be determined using quantitative PCR (qPCR) where load of pathogen is determined using threshold cycle ($C_t$) values[8–10]. However, studies have shown that the severity of AGD established using Taylor scoring system vs. $C_t$ values using qPCR have high genetic (close to unity) and phenotypic (0.81) correlations revealing that both are potentially the same trait[8,10].

Amoebic gill disease has been a major problem over many years in farmed Atlantic salmon (*Salmo salar*) of Tasmania, and cleaning of amoeba from the gills requires freshwater treatments which costs ~10–20% of the total production cost[4]. The treatment is also a welfare issue for the fish due to the physiological stress caused by freshwater bath.

In Northern Atlantic, the presence of *P. perurans* has been documented for more than a decade, but for a long time the cold water seemed to have prevented an epidemic of AGD[3]. However, warm and dry weather conditions in 2011 and 2012 for Ireland and Scotland, and later in 2012–2013, at Northern Isles (Orkney and Shetland), Norway and the Faroe Islands caused major AGD outbreaks on farmed Atlantic salmon[11], and AGD became the largest infectious health problem for the salmon industry in Ireland, Scotland and France those years[12].

AGD is a rising threat for Norwegian salmon with first documented occurrence in 2006[13], and since then amoeba has been regularly reported every year on the southwest coast and further north[14] in Norway.

[1]Department of Breeding and Genetics, Nofima AS, P.O. Box 210, N-1431, Ås, Norway. [2]Marine Harvest ASA (old name) with new name Mowi Genetics AS, 5035, Bergen, Norway. *email: luqman.aslam@nofima.no

Norwegian Atlantic salmon populations from the two breeding companies (Marine Harvest ASA and SalmoBreed AS) have shown genetic variation for resistance against AGD both in field ($h^2$ of 0.12–0.20) and challenge test ($h^2$ of 0.09–0.13) conditions[15]. However, reported heritability estimates for AGD score in Tasmanian population showed higher range with estimates of 0.10 to 0.48[7,8,15–17], with lower heritability estimates obtained from the first infection and the higher estimates for the subsequent infections. Tasmanian research has shown that the resistance against first and later subsequent infections are different traits with poor genetic correlations (average $r_g = 0.24$)[16]. Selective breeding has been effective to increase the intervals between two consecutive baths/treatments which lead to overall reduction in number of baths/treatments and ultimately reduction in the expenses incurred on AGD[16]. However, addition of AGD resistance in breeding goal traits will reduce selection response for other traits, particularly when AGD show unfavourable genetic correlations to any other traits[18,19]. The use of marker assisted selection (MAS) and/or genomic selection (GS) using molecular markers that are directly or indirectly linked to variation in causal loci could provide potent tools to overcome these challenges which may increase both selection accuracy as well as selection intensity as this allow also for within family selection[20,21]. In addition, identification and subsequent fine-mapping of QTL regions should allow for the pinpointing of genes that underlie such traits. Significant association between genetic markers and quantitative traits of economic importance have been reported in Atlantic salmon[22–24].

A few studies have been conducted on detection of QTL for resistance against AGD, where results have shown that the genetic architecture for AGD resistance trait is polygenic in nature[8,9]. Moreover, the results on transcriptomic profiles of Atlantic salmon in response to AGD infection seems to also explain polygenic nature of this trait with changes in expression of many genes in infected individuals[9,25]. The genes with functional properties in the immune system (e.g. interleukin-1 beta, a pro-inflammatory cytokine)[9,26,27] as well as in cellular-adhesion (e.g. CCAAT/enhancer binding protein beta)[9,25] were reported to be of importance for playing role in trait variation.

Availability and popularity of advanced GS by which breeding values of individuals are predicted using statistical methods has become a method of choice in recent era. The GS involves genotypic data for genome-wide distributed single nucleotide polymorphism (SNP) markers[28] and provides opportunity to rank individuals within and across families. The feasibility of GS depends on the availability of a high-quality SNP genotyping platform and on extensive trait records collected in the reference populations. It has already been a widely used approach in many livestock and aquaculture species[29–33] due to relative reduction in genotypic and sequencing costs which is primarily applied for the improvement of traits of economic and welfare importance (e.g. disease resistance).

The aim of the current study was to identify the genetic basis of host resistance to AGD by performing genome-wide association analysis (GWAS), and compare accuracies of genomic vs. pedigree-based predictions.

## Materials and Methods

**Resource population.** The population used originated from Marine Harvest (MH) breeding nucleus in Ireland which was developed with a cross of parents mated in 1:2 male to female ratio. The starting population had 150 full sib families with 40 full-sibs per family. Families were communally reared from the eyed egg stage, and the tagging was performed at an average body weight of ~45 g. At tagging, fin clip samples were collected and preserved in 100% ethanol for further DNA extraction and genotyping (~50 K SNP chip). Pedigree was constructed using a panel of 65 SNP markers.

**Field test.** A population of 6,100 fish at an average weight of 61 g were placed in a sea net-cage on 24. April 2014 at the South West farm of Marine Harvest Ireland, where AGD is a common problem. Fish population was allowed to develop AGD from the natural concentration of amoeba in sea. In June 2014, AGD was reported at the farm holding the fish and presence of *P. perurans* was confirmed by PCR. The monitoring for the development of AGD in the test cage was done by regular gill-scoring of a small number (~10–15) of fish per week. Gills were scored from 0 to 5 as described by Taylor *et. al.*[7] with intensity of infection increasing with ascending order of score level. Major scoring for AGD phenotypes (first infection) was done on 29$^{th}$ July 2014 when the estimated average gill-score in the sea net-cage passed score 2. Fish were haphazardly picked from the cage and uniquely gill scored by one of the three scorers (A, B and C). The sampling continued until the daylight was no longer appropriate for gill scoring which resulted in 3,663 gill-scored fish. The day after the fish were treated with fresh water to kill the amoeba, after which the fish were allowed to develop another round of a natural AGD infection (second infection). Scoring for second infection was also done by the same three scorers on 12$^{th}$ of September 2014, which resulted in 3,511 gill-scored fish.

**DNA extraction and genotyping.** From the 3,663 gill-scored fish during first infection, 1,190 fish (10 sibs per family) were randomly selected from 119 full-sib families for further processing. Genomic DNA was extracted from the fin clips using a commercial kit (DNeasy Blood & Tissue Kit, Qiagen), following the manufacturer's instructions. Fish were genotyped using a 57 K axiom Affymetrix SNP Genotyping Array (NOFSAL2). After genotyping, we were left with a total of 1334 (193 Parents + 1141 Progeny) individuals from 119 full-sib families of 6–10 offspring per family and the rest either failed to genotype or filtered out during genotype calls due to the poor genotype quality.

**Filtering of SNPs.** Genotypic data was filtered using the Plink software[34], and SNPs were excluded with minor allele frequency (MAF) lower than 5%, missing rate higher than 15%, and Mendelian errors. Finally, approximately 54 K SNPs were retained for analyses.

**Statistical analysis.** Analyses were performed for both first and second infection but the second infection did not show any genetic variation for this dataset of 1,141 phenotypes. Hence, following analyses were continued with gill-scores obtained in first infection only.

The applied model also included a fixed effect of scorer and an additional random effect common to full-sibs other than additive genetics. However, both effects were found to be non-significant ($P > 0.05$) and was therefore omitted from the model.

*Genome wide association analysis (GWAS).* Genome wide association analysis was performed using the following linear mixed animal model implemented in GCTA program with "–mlma-loco" and "–reml" functions[35].

$$y = \mu + X\alpha + Zu + e$$

where $y$ is a vector of n (n = 1,141) AGD scores, μ is an overall mean; $X$ is the incidence matrix for SNP containing marker genotypes coded as $0 = AA$, $1 = AB|BA$, $2 = BB$, $\alpha$ is the allele substitution effect of each SNP, $Z$ is the incidence matrix of genotyped individuals, $u$ is the vector of genomic breeding values with $u \sim N(0, G\sigma_u^2)$, where $\sigma_u^2$ is the additive genetic variance, and $e$ is the vector of random residual effects with $e \sim N(0, I\sigma_e^2)$. The G matrix is a genomic relationship matrix (GRM) which was computed according to VanRaden (2008)[36] as $\frac{ZZ'}{2 * \sum_{i=1}^{Nsnp} p_i(1-p_i)}$; where $p_i$ is the allele frequency of second allele and $Nsnp$ is the total number of SNP markers.

SNPs were considered genome wide significant when they exceeded the Bonferroni threshold[37] for multiple testing (alpha = 0.05) of $0.05/tg$, where $tg$ =53,865 (total number of SNPs genome-wide) and graded as chromosome-wide significant when Bonferroni threshold for multiple testing surpassed (alpha = 0.05) $0.05/tc$, where $tc$ =1,796 (average number of SNPs per chromosome). Genome-wide significant threshold used in this study was considered to be $P \leq 9.28 \times 10^{-7}$ which is equivalent to $-log_{10}(P) = 6.03$, while chromosome-wide significant threshold was opted to be $P \leq 2.78 \times 10^{-5}$ which is equal to $-log_{10}(P) = 4.55$

Quantile-quantile (q-q plot) plot with distribution of observed vs. expected p-values was checked, and the Inflation factor (lambda, λ) was calculated using following equation

$$lambda(\lambda) = \frac{median(\chi^2)}{0.456}$$

*Estimation of SNP variances.* Variances explained by the top significant SNP(s) were estimated using following two approaches (direct and indirect).

For the direct approach, variances explained by the top significant SNP(s) were estimated as $=2p_iq_i\alpha_i^2$ (Falconer and Mackay (1996)[38]). Therefore, the proportion of the of genetic ($\%varG_{SNP_i}$) or phenotypic ($\%varP_{SNP_i}$) variances captured by these markers equals $\frac{var_{SNP_i}}{\sigma_g^2} \times 100$ and $\frac{var_{SNP_i}}{\sigma_p^2} \times 100$, respectively. Where, $p_i$ and $q_i$ are allele frequencies for the major and the minor alleles respectively, whereas $\sigma_g^2$ and $\sigma_p^2$ are the genetic and phenotypic variances computed with the above animal model using genomic relationship matrix.

For the indirect approach, the proportion of the genetic or phenotypic variance explained by the genome-wide significant SNP(s) was estimated using the model: $y = \mu + GWS + Zu + e$. Where, $GWS$ are the genome wide significant SNP(s), the $G$ matrix used in this model was constructed with all other SNPs except genome-wide significant SNP ($GWS$). The variance (genetic or phenotypic) explained by the $GWS$ SNPs was expressed as a reduction in the total genetic or phenotypic variance.

*Breeding value estimation.* Pedigree as well as genomic breeding values (PEBVs vs. GEBVs) were computed using full (n=3,663) or reduced (n=1,141) datasets. The full dataset contained phenotypic records on all the recorded animals (n=3,663), while the reduced dataset (n=1,141) included phenotypic records on only the genotyped individuals which is a subset of the full data. Breeding values were estimated by applying the same model as described under the "GWAS" section of materials and methods, except that the marker effect ($X\alpha$) was excluded from the model and the genomic relationships ($G$) was constructed using all SNPs that passed quality control. Breeding values for all scenarios were computed using ASreml v4.0[39] program. Pedigree-based breeding values were computed by replacing the G matrix with the numerator relationship matrix (A). Pedigree breeding values were obtained with the dataset consisting of phenotypic records from only the genotyped (PBLUP_I) or from all the phenotyped (PBLUP_II) animals. Similarly, genomic breeding values were computed using records from only the genotyped animals (GBLUP) or a combined relationship matrix that uses all genotyped and phenotyped (ssGBLUP) animals. Whereas the G matrix was used for the GBLUP analysis, the realized relationship matrix (H) replaces G. The inverse of the H matrix (Legarra *et al*., 2009; Misztal *et al*., 2009) was constructed as follows:

$$H^{-1} = A^{-1} + \begin{bmatrix} 0 & 0 \\ 0 & (0.95G + 0.05A_{22})^{-1} - A_{22}^{-1} \end{bmatrix}$$

where $G$ is as described above and $A_{22}$ is the pedigree-based relationship matrix for genotyped animals. The variance components ($\sigma_u^2 = 0.120$ and $\sigma_e^2 = 0.480$) used for the genomic prediction analysis were computed from the full dataset[15] and was fixed in all analysis.

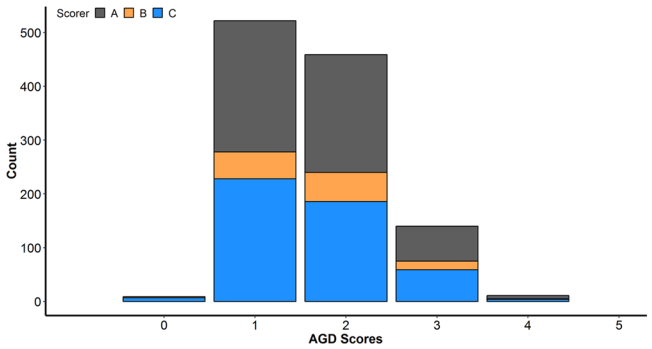


**Figure 1.** Distribution of AGD gill-scores and frequency of scoring by each scorer.

| Model | $\sigma_g^2$ | $\sigma_e^2$ | $\sigma_p^2$ | Genomic $h^2$ |
|---|---|---|---|---|
| 1 | 0.061 (0.021) | 0.472 (0.025) | 0.533 (0.023) | 0.114 (0.037) |
| 2 | 0.036 (0.018) | 0.467(0.025) | 0.503 (0.023) | 0.071 (0.036) |

**Table 1.** Estimates of variance components and heritability with standard errors (in parenthesis) using the genomic relationship matrix. $\sigma_g^2$ = Genetic variance; $\sigma_p^2$ = Phenotypic variance; $\sigma_e^2$ = Residual variance; $h^2$ = Heritability; 1 = Model without any SNP used as fixed effect; 2 = Model with genome-wide significant SNP used as fixed effect.

| SNP-ID | Ssa | Pos(bp) | A1 | A2 | MAF | α | SE | P | *VarP* | *varG* (%) | *varP* (%) |
|---|---|---|---|---|---|---|---|---|---|---|---|
| AX-88266207 | 02 | 1967633 | A | B | 0.165 | −0.265 | 0.052 | 4.12E-07 | 0.019 | 32.6 | 3.63 |
| AX-87970438 | 05 | 76572428 | A | B | 0.184 | −0.228 | 0.050 | 5.73E-06 | 0.016 | 26.3 | 2.92 |
| AX-88137791 | 02 | 2059898 | B | A | 0.184 | −0.208 | 0.047 | 7.96E-06 | 0.013 | 22.0 | 2.45 |
| AX-87975635 | 01 | 52913027 | A | B | 0.314 | 0.159 | 0.037 | 2.08E-05 | 0.011 | 18.4 | 2.05 |
| AX-87017245 | 01 | 53002456 | A | B | 0.313 | 0.159 | 0.037 | 2.13E-05 | 0.011 | 18.3 | 2.04 |

**Table 2.** The top five significant SNPs detected in GWAS analysis ranked with respect to level of significance. Ssa = *Salmo salar* chromsomes; Pos(bp) = Physical position of SNP; A1 & A2 = Minor & major alleles, respectively; MAF = Minor allele frequency; α = Allele substitution effect; SE = Standard error; P = Significance value; *varP* = Phenotypic variance explained; *varG* (%) = Proportion of genotypic variance explained; *varP*(%) = Proportion of phenotypic variance explained.

*Cross-validation and accuracy of prediction.* Within family cross-validation scheme was used to assess the accuracy of the predicted breeding values. The phenotypes of four offspring per sire family were randomly masked as validation dataset (296 offspring out of 1,141) and the remaining animals were used as training dataset. This procedure was replicated 50 times and for each replicate, accuracy of predictions were computed as:

the correlation ($r_{corr}$) of either the estimated pedigree (PEBV) or the genomic (GEBV) breeding value with the pre-corrected phenotype $y_{adj}$ which was scaled by the square root of heritability as $r_{corr} = \frac{\rho\left(G[P]EBV, y_{adj}\right)}{\sqrt{h^2}}$; where $\rho$ = correlation coefficient, $G[P]EBV$ = breeding values estimated using genomic (GBLUP) or pedigree (PBLUP) information; $h^2 = 0.20 \pm 0.03$ as reported by Lillehammer *et al.* 2019[15].

## Results

**Descriptive statistics.** The 1,141 gill-scored fish had a mean gill-score of 1.67 and standard deviation of 0.73. The distribution of the gill-scores and number of individuals scored by each scorer is given in Fig. 1. The number of fish scored by each person was 123, 484, and 534 (B, C and A, respectively), and the gill-scores ranged from 0–4 with very low frequency of the extreme phenotypes (gill-scores 0 and 4).

**Genomics based variance components.** Estimated variance components using narrated models with and without the genome-wide significant SNP are given in Table 1. Genomic heritability for resistance to AGD was 11.4%, but was reduced to 7.1%, or by 37.7%, when the genome-wide significant SNP (see Table 2) were accounted for in the model.

**Genome wide association analysis and SNP variances.** GWAS analysis using 53,865 SNPs on 1,141 recorded individuals resulted in a total of 5 SNPs which crossed genome or chromosome-wide significant level (Fig. 2). These five most significant SNPs are distributed across three different chromosomes 1, 2, and 5 (Ssa01,

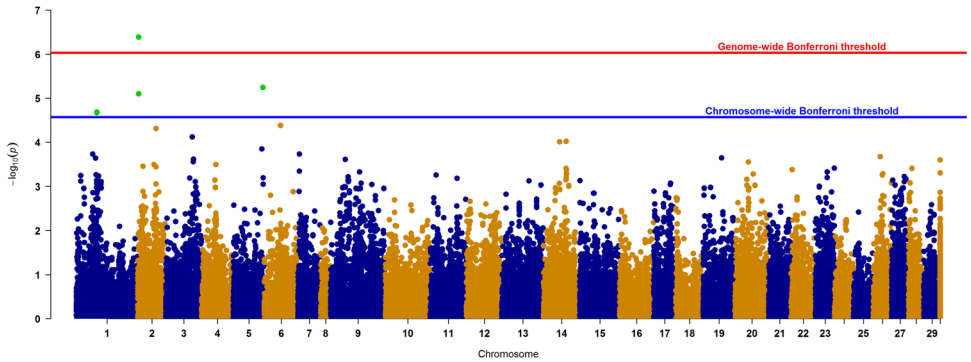


**Figure 2.** Manhattan plot of GWAS with p-values distributed across different chromosomes. Markers crossing genome and/or chromosome wide Bonferroni threshold are dotted in green color. Chromosome 30 represent markers belonging to unknown chromosome(s).

| Ssa | SNP | Annotation | Genes (~ ± 20 Kb) | Description | Ref |
|---|---|---|---|---|---|
| 02 | AX-88266207 | Intronic | *tnxb.b* | This gene encodes a member of the tenascin family of extracellular matrix glycoproteins. This protein is thought to function in matrix maturation during wound healing. | 41 |
| | | | *slc45a4* | Solute carrier family 45, member 4, found to be functioning as sucrose transporter | 54 |
| | | | *c4* | It is central to the activation of both the classical and the lectin pathways of complement activation. The complement system is a part of the immune system that enhances (complements) the ability of antibodies and phagocytic cells to clear pathogens from an organism. | 45 |
| 05 | AX-87970438 | Intronic | *slc44a4* | Solute carrier family 44 member 4 involved in the uptake of choline by cholinergic neurons. | 41 |
| | | | *ano10a* | The transmembrane protein encoded by this gene is a member of a family of calcium-activated chloride channels. | 41 |
| 01 | AX-87017245 | 3'UTR | *kcng3* | Potassium voltage-gated channel modifier subfamily G member 3. It is diverse in functionality includes epithelial electrolyte transport, heart rate etc. | 55 |

**Table 3.** Summary for the functions of important candidate genes at QTL region. Genes are searched within approximately ± 20Kb region of the significant SNPs, and detected candidate genes might be playing role in variation for AGD phenotype.

Ssa02 and Ssa05); with 2, 2 and 1 SNPs respectively (Table 2). The λ value (i.e. magnitude of the deviation inflation/deflation of p-values) for the analysis was recorded to be 1.094, and distribution of p-values are presented in a Q-Q plot (Supplementary Fig. S1.1).

Phenotypic variance and the proportion of genetic and phenotypic variances explained by each of the five most significant SNPs are shown in Table 2. The proportion of the genetic variation (estimated with the direct approach) captured by each of these SNPs ranged from 18.3–32.6%. While the proportion of genetic variance explained by the genome-wide significant SNP estimated with the indirect method was 37.7%.

Potential genes with effect on AGD were searched around ±20 Kb region of the genome or chromosome-wide significant SNPs using salmon genome database[40]. The highest significant SNP of chromosome 2 was annotated as an intronic SNP located within the intron of "*tnxb.b*" gene and another gene "*c4*" is located upstream to this SNP (Table 3, Fig. S1.2). Both genes (*tnxb.b* and *c4*) are known to play a role in the immune system[41].

**Accuracy of prediction.** Pedigree and genomic-based prediction accuracies for amoebic gill disease are given in Table 4. Overall, genomic (GBLUP and ssGBLUP) based accuracies ($r_{cor}$) were higher compared to using pedigree information (PBLUP_I and PBLUP_II). Accuracies from the model that used phenotypes (n=1,141) of only genotyped individuals (PBLUP_I and GBLUP) was much lower than using phenotype information (n = 3,663) from all individuals (PBLUP_II and ssGBLUP) that were recorded (Table 4). The increase in accuracy ($r_{cor}$) when the number of phenotypic records tripled (n = 1141 vs n = 3,663) was about 34%. The blending of genomic and pedigree information led to about 12% increase in accuracy (ssGBLUP =0.47 vs PBLUP_II = 0.43).

**Ethical approval for the use of animals in this study.** Although animals were used in this work, no direct experiments were performed on them. Permission was taken to use the site (InishFanard, on the south west coast of Ireland) for farming and production, and fish were moved to the site after obtaining the approval. Health status of fish was monitored regularly, and in-case of any outbreak fish were recorded and treated as part of routine practices. The studied population faced natural outbreak of AGD twice, and the recording of phenotype(s) is a routine procedure of breeding companies. The collection of tissue samples was carried out by highly skilled and experienced personnel from the breeding company. Tagging of fish, and the sampling the fin clips is

| Model | Number of fish | | | Accuracy |
|---|---|---|---|---|
| | Total | Training | validation | |
| PBLUP_I | 1,141 | 845 | 296 | $0.32_{(0.10)}$ |
| GBLUP | 1,141 | 845 | 296 | $0.35_{(0.09)}$ |
| PBLUP_II | 3,663 | 3,367 | 296 | $0.43_{(0.11)}$ |
| ssGBLUP | 3,663 | 3,367 | 296 | $0.47_{(0.12)}$ |

**Table 4.** Genomic vs. pedigree based prediction accuracies for amoebic gill disease. PBLUP_I and PBLUP_II – Pedigree based breeding values using phenotypes from only genotyped or all phenotyped animals, respectively. GBLUP - Genomic breeding values from only genotype animals, and ssGBLUP - Genomic breeding values from all genotyped and phenotyped animals obtained with a combined relationship matrix (H).

not considered as an experimental intervention in the EU. Hence, no approval from the ethics committee was necessary according to local legislation.

## Discussion

Amoebic gill disease (AGD) is an increasing threat for the Atlantic salmon industry and understanding genetic basis of AGD and the application of advanced selection methods could lead to robust and sustainable salmon production. Current study aimed at detection of QTL(s) for AGD resistance as well as determine consistency in accuracy of genomic vs. pedigree based prediction methods.

The model applied in our statistical analyses did not include scorers as fixed effect because it was found to be non- significant ($P > 0.05$) for this dataset. However, scoring system for determining the severity of AGD is subjective, and it is therefore recommended to include scorer-effect in the model as it might have a significant effect as seen in some other experiments[15]. The observed genomic heritability of ~11.5% (for AGD score at first infection) obtained from our analysis was found to be lower than the previously reported estimates by Lillehammer *et al.*[15] with 20% and 11% in field and challenge conditions respectively, and a field test based heritability of 14% (first infection) reported by Kube *et al.*[16]. The reported heritability estimates on AGD scores recorded at the third infection in challenge conditions fall within medium to higher level with estimates ranging from 0.24 to 0.48[7,8]. The difference in estimates could have been due to multiple reasons e.g. difference in adopted methodology for the estimations, population differences, infection types (first vs. subsequent later infections) used as phenotype which might activate different immune responses, e.g. first infection should activate innate while subsequent later infection should be mainly pursued by acquired immune system, and/or the total number of markers used in this study perhaps could not explain the total genetic variance.

The detected significant QTLs at chromosomes Ssa01, Ssa02 and Ssa05 for the AGD scores after first infection-did not show concordance with the detected QTLs in previous studies of Robledo *et al.* and Boison *et al.*[8,9]. This disagreement could likely be due to the differences in infection conditions (challenge test in both studies vs natural field outbreak in current study), infection type or time of recording (first vs. subsequent infections), and/or differences in populations. The studied populations in both Robledo *et al.* and Boison *et al.*[8,9] were challenge tested and had a trait in common where amoebic load was recorded using qPCR based $C_t$ values. The QTL results from Robledo *et al.* and Boison *et al.*[8,9] also did not show concordance in any of the detected QTLs which further highlights the complexity of this trait, and perhaps indicates the cruciality of factors i.e. environment, time and type of recordings and genetic background of populations.

The top three most significant SNPs of GWAS analysis showed a favorable effect on the trait with negative $\alpha$ values which represent a reduction in AGD score (Table 2), while significant SNPs on Ssa01 had unfavorable effect on the trait (Figure S1.4, Table S1.1) with positive $\alpha$ values. Individual SNP specific genetic variances estimated using direct method (Table 2) cannot be added to find their cumulative effect because the top 3 SNPs on Ssa02 and Ssa05 are in high linkage disequilibrium (LD) with LD values ranging from 0.77–1.0, and the two SNPs on Ssa01 are also in complete LD of 1.0 (Supplementary Table S1.1 & Figure S1.4). Rather an average of the individual SNP genetic variances is likely to provide a better estimate which means that the top 3 SNPs and the significant SNPs on Ssa01 account for 26.9% and 18.3% of the genetic variance for resistance to AGD, respectively. However, as the indirect method (SNP as a fixed effect in the model) yielded a higher genetic variance estimate (~37%) which indicates that the individual SNPs are not completely linked as also indicated by the LD values less than unity. Application of indirect method with fixing the genome-wide significant SNP in the model appeared to cause insignificance of previously chromosome-wide significant SNPs of Ssa02 and Ssa05, and also shrinkage in p-values for all the other SNPs (Supplementary Figure S1.7). High shrinkage of p-values for "AX-87970438" and "AX-88137791" SNPs can be explained due to the existence of high co-linearity among the top 3 SNPs. Moreover, this high observed shrinkage in p-values of "AX-87970438" and "AX-88137791" SNPs also justifies the above described averaging function in-case of direct method instead of additive when estimating variances. The estimatedvariances explained by these significant (genome-wide and/or chromosome-wide) SNPs are likely to be inflated and could be due to the Beavis effect[42]. Large impact of these SNPs on genetic variation does not necessarily mean that the SNPs are causative mutations, but that these SNPs explain an important amount of the QTL variation, either directly or through LD with the causative mutations.

It is interesting to note that the significant SNPs on Ssa02 are located within ~92 Kb region (Figs. S1.2–1.4) and show high LD of 0.77, but the SNPs "AX-88137791" and "AX-87970438" located on chromosome Ssa02 and Ssa05, respectively showed complete LD of 1.0. The LD information among the significant SNPs in Figure S1.4 and Table S1.1 gives strong impression of co-segregation pattern for the SNPs on Ssa02 and Ssa05. We checked if SNPs were belonging to the homeologous regions of Ssa02 and Ssa05[40], and interestingly they were positioned

in homeologous block, 2p-5q_1 of both chromosomes (Ssa02 and Ssa05). However, when their positions were checked towards a recently developed linkage map (unpublished data), all three significant SNPs of Ssa02 and Ssa05 were found to be closely located on the same chromosome Ssa02 within a distance of 0.62 cM. The LD pattern and linkage map suggest that the top significant SNPs of Ssa02 and Ssa05 are located at the same chromosome which is discordant with the physical map and could be due to assembly error or complicated long-range linkage disequilibria explained by Koch *et al.*[43]. Manhattan plot for all chromosomes and the distribution of P-values for the SNPs on Ssa02 were replotted after correcting positions of significant markers (Figure S1.5 and S1.6) which provided relatively better shape of the QTL peak.

The region of ±20 Kb surrounding significant SNPs on Ssa02 include genes *tnxb.b* and *c4* which have been reported to be involved in the immune response (Table 3), and may play a role in AGD resistance. One of the SNPs on Ssa02 (the highest significant SNP of this study) is located within the intron of *tnxb.b* and this gene is reported to produce extracellular matrix glycoproteins which has anti-adhesive effects and is thought to function in matrix maturation during wound healing[44]. The *C4* gene is sandwiched between the two mutations located downstream (~38Kb distant) of the highest and upstream (~71.5 Kb distant) to the second highest SNPs of Ssa02. Gene *C4* is known to be central to the activation of both the classical and the lectin pathways of complement system that enhances the ability of antibodies and phagocytic cells to clear pathogens from an organism[45]. The functional properties of these candidate genes (*tnxb.b & C4*) with their role in immune system and in cellular-adhesion agrees with transcriptomics results of previous studies where immune and cellular-adhesion functionality genes were detected to be differentially expressed[9,25]. Available annotated information on both the genes (*tnxb.b & C4*) signifies the impact with strong signal that the QTL of Ssa02 could directly and/or indirectly be linked with variation in expression level and/or function of these gene(s), which ultimately cause variation in AGD score. However, further studies on QTL validations as well as advanced assays like *in situ* hybridization technique to detect localization of gene expression, differential expression or sequencing of selected genes in susceptible and resistance fish might lead to a better understanding of the biological mechanism/pathways in response to this pathogen.

Overall, we observed 9.3% increase in accuracy with genomic information depending on the method used to calculate the accuracies. Similar trend of higher accuracies using genomic vs. pedigree-based information has been reported in Atlantic salmon[46,47] for parasite and pathogen resistance traits, as well as for production traits in livestock species[48–50]. The advantage of GBLUP over PBLUP is because realized genomic-based relatedness between animals deviate from pedigree-based relationship coefficients.

Our results showed that single-step methodology (ssGBLUP), which take advantage of pedigree, phenotypic and genomic information simultaneously, gave higher prediction accuracies compared to PBLUP and/or GBLUP, which used only pedigree or genomic information. Similar results obtained with ssGBLUP were reported in salmonids[51,52] and cattle[53].

## Conclusion

A SNP array based genotyping of a population of Atlantic salmon which was field recorded for resistance to AGD revealed one genome-wide significant and the two suggestive QTLs distributed over chromosome 1, 2 and 5. Three candidate genes; *c4*, *tnxb*, and *slc44a4* were found within nearly 20 Kb flanking region of the detected loci on chromosome 2. The *tnxb and c4* genes are known to be a part of innate immune system, which may be involved in resistance to AGD. Genomic prediction using SNP based genotypic data improved prediction accracy with 9–17% over the pedigree-based predictions which highlights both the potential and importance of genomic selection in commercial breeding programs.

## Data availability

Most of the data supporting these findings are contained within the manuscript. Specific queries regarding data can be made available upon request through corresponding author.

Received: 7 October 2019; Accepted: 30 March 2020;
Published online: 15 April 2020

## References

1. Crosbie, P. B. B., Bridle, A. R., Cadoret, K. & Nowak, B. F. *In vitro* cultured Neoparamoeba perurans causes amoebic gill disease in Atlantic salmon and fulfils Koch's postulates. *International Journal for Parasitology* **42**, 511–515, https://doi.org/10.1016/j.ijpara.2012.04.002 (2012).
2. Young, N. D., Dyková, I., Snekvik, K., Nowak, B. F. & Morrison, R. N. Neoparamoeba perurans is a cosmopolitan aetiological agent of amoebic gill disease. *Diseases of Aquatic Organisms* **78**, 217–223 (2008).
3. Munday, B. L., Zilberg, D. & Findlay, V. Gill disease of marine fish caused by infection with Neoparamoeba pemaquidensis. *Journal of Fish Diseases* **24**, 497–507 (2001).
4. B. F. Nowak. (eds P.T.K. Woo & K. Buchmann) 1-18 (CAB Interanational, Oxfordshier, UK., 2012).
5. Adams, M. B. & Nowak, B. F. Distribution and structure of lesions in the gills of Atlantic salmon, *Salmo salar L.*, affected with amoebic gill disease. *J. Fish Dis.* **9**, 535–542 (2001).
6. Clark, A. & Nowak, B. F. Field investigations of amoebic gill disease in Atlantic salmon, *Salmo salar* L., in Tasmania. *Journal of Fish Diseases* **22**, 433–443 (1999).
7. Taylor, R. S., Muller, W. J., Cook, M. T., Kube, P. D. & Elliott, N. G. Gill observations in Atlantic salmon (*Salmo salar, L.*) during repeated amoebic gill disease (AGD) field exposure and survival challenge. *Aquaculture* **290**, 1–8 (2009).
8. Robledo, D., Matika, O., Hamilton, A. & Houston, R. D. Genome-Wide Association and Genomic Selection for Resistance to Amoebic Gill Disease in Atlantic Salmon. *G3 (Bethesda, Md.)* **8**, 1195–1203, https://doi.org/10.1534/g3.118.200075 (2018).
9. Boison, S. A., Gjerde, B., Hillestad, B., Makvandi-Nejad, S. & Moghadam, H. Genomic and transcriptomic analysis of amoebic gill disease resistance in Atlantic salmon (*Salmo salar L.*). *Front. Genet.*, https://doi.org/10.3389/fgene.2019.00068 (2019).
10. Gjerde, B. *et al.* Estimates of genetic correlations between susceptibility of Atlantic salmon to amoebic gill disease in a bath challenge test and a field test. *Aquaculture* **511**, 734265, https://doi.org/10.1016/j.aquaculture.2019.734265 (2019).

11. Hjeltnes, B. *et al*. Risk assessment of amoebic gill disease. 39 (Oslo, Norway, 2014).
12. Mitchell, S. O. & Rodger, H. D. A review of infectious gill disease in marine salmonid fish. *Journal of Fish Diseases* **34**, 411–432, https://doi.org/10.1111/j.1365-2761.2011.01251.x (2011).
13. Steinum, T. *et al*. First cases of amoebic gill disease (AGD) in Norwegian seawater farmed Atlantic salmon, *Salmo salar L*., and phylogeny of the causative amoeba using 18S cDNA sequences. *Journal of Fish Diseases* **31**, 205–214 (2008).
14. Hjeltnes, B. *et al*. Panel on Animal Health and Welfare; Risk assessment of amoebic gill disease. 39 (Norwegian Scientific Committee for Food Safety (VKM), Oslo, Norway, 2014).
15. Lillehammer, M. *et al*. Genetic parameters of resistance to amoebic gill disease in two Norwegian Atlantic salmon populations. *Aquaculture* **508**, 83–89 (2019).
16. Kube, P. D., Taylor, R. S. & Elliott, N. G. Genetic variation in parasite resistance of Atlantic salmon to amoebic gill disease over multiple infections. *Aquaculture* **364**, 165–172, https://doi.org/10.1016/j.aquaculture.2012.08.026 (2012).
17. Taylor, R. S., Wynne, J. W., Kube, P. D. & Elliott, N. G. Genetic variation of resistance to amoebic gill disease in Atlantic salmon (*Salmo salar*) assessed in a challenge system. *Aquaculture*, S94-S99 (2007).
18. Haffray, P. *et al*. Negative genetic correlations between production traits and head or bony tissues in large all-female rainbow trout (*Oncorhynchus mykiss*). *Aquaculture* **368–369**, 145–152, https://doi.org/10.1016/j.aquaculture.2012.09.023 (2012).
19. Yáñez, J. M. *et al*. Negative genetic correlation between resistance against Piscirickettsia salmonis and harvest weight in coho salmon (Oncorhynchus kisutch). *Aquaculture* **459**, 8–13, https://doi.org/10.1016/j.aquaculture.2016.03.020 (2016).
20. Meuwissen, T. H. E. & Goddard, M. E. The use of marker-haplotypes in animal breeding schemes. *Genetics Selection Evolution* **28**, 161–176 (1996).
21. Pyasatian, N., Fernando, R. L. & Dekkers, J. C. M. Genomic selection for marker-assisted improvement in line crosses. *Theor Appl Genet* **115**, 665–674 (2007).
22. Everett, M. V. & Seeb, J. E. Detection and mapping of QTL for temperature tolerance and body size in Chinook salmon (Oncorhynchus tshawytscha) using genotyping by sequencing. *Evolutionary Applications* **7**, 480–492, https://doi.org/10.1111/eva.12147 (2014).
23. Gonen, S. *et al*. Mapping and validation of a major QTL affecting resistance to pancreas disease (salmonid alphavirus) in Atlantic salmon (Salmo salar). *Heredity* **115**, 405–414, https://doi.org/10.1038/hdy.2015.37 (2015).
24. Moen, T. *et al*. Epithelial Cadherin Determines Resistance to Infectious Pancreatic Necrosis Virus in Atlantic Salmon. *Genetics* **200**, 1313–1326 (2015).
25. Wynne, J. W. *et al*. Resistance to amoebic gill disease (AGD) is characterised by the transcriptional dysregulation of immune and cell cycle pathways. *Developmental & Comparative Immunology* **32**, 1539–1560, https://doi.org/10.1016/j.dci.2008.05.013 (2008).
26. Bridle, A. R., Morrison, R. N., Cupit Cunningham, P. M. & Nowak, B. F. Quantitation of immune response gene expression and cellular localisation of interleukin-1β mRNA in Atlantic salmon, *Salmo salar L*., affected by amoebic gill disease (AGD). *Veterinary Immunology and Immunopathology* **114**, 121–134, https://doi.org/10.1016/j.vetimm.2006.08.002 (2006).
27. Morrison, R. N. *et al*. Molecular cloning and expression analysis of tumour necrosis factor-α in amoebic gill disease (AGD)-affected Atlantic salmon (*Salmo salar L*.). *Fish & Shellfish Immunology* **23**, 1015–1031, https://doi.org/10.1016/j.fsi.2007.04.003 (2007).
28. Meuwissen, T. H., Hayes, B. J. & Goddard, M. E. Prediction of total genetic value using genome-wide dense marker maps. *Genetics* **157**, 1819–1829 (2001).
29. Fulton, J. E. Genomic selection for poultry breeding. *Animal Frontiers* **2**, 30–36, https://doi.org/10.2527/af.2011-0028 (2012).
30. Gutierrez, A. P., Yáñez, J. M., Fukui, S., Swift, B. & Davidson, W. S. Genome-wide association study (GWAS) for growth rate and age at sexual maturation in Atlantic salmon (*Salmo salar*). *PLoS One*, **10**, e0119730 (2015).
31. Marle-Köster, E. v., Visser, C. & Berry, D. P. A review of genomic selection - Implications for the South African beef and dairy cattle industries. *S. Afr. J. Anim. Sci.* **43** (2013).
32. Schefers, J. M. & Weigel, K. A. Genomic selection in dairy cattle: Integration of DNA testing into breeding programs. *Animal Frontiers* **2**, 4–9 (2012).
33. Yáñez, J. M., Newman, S. & Houston, R. D. Genomics in aquaculture to better understand species biology and accelerate genetic progress. *Frontiers in Genetics* **6**, 128, https://doi.org/10.3389/fgene.2015.00128 (2015).
34. Purcell, S. *et al*. PLINK: A Tool Set for Whole-Genome Association and Population-Based Linkage Analyses. *The American Journal of Human Genetics* **81**, 559–575, https://doi.org/10.1086/519795 (2007).
35. Yang, J., Lee, S. H., Goddard, M. E. & Visscher, P. M. GCTA: A tool for genome-wide complex trait analysis. *Am. J. Hum. Genet.* **88**, 76–82, https://doi.org/10.1016/j.ajhg.2010.11.011 (2011).
36. VanRaden, P. M. Efficient methods to compute genomic predictions. *J. Dairy Sci.* **91**, 4414–4423, https://doi.org/10.3168/jds.2007-0980 (2008).
37. Bonferroni, C. E. *Il calcolo delle assicurazioni su gruppi di teste*. (In Studi in Onore del Professore Salvatore Ortu Carboni, 1935).
38. Hill, W. G. & Mackay, T. F. C. D. S. Falconer and Introduction to Quantitative Genetics. *Genetics* **167**, 1529–1536 (2004).
39. ASReml user guide release 3.0 (VSN International Ltd, Hemel Hempstead, UK, 2009).
40. Lien, S. *et al*. The Atlantic salmon genome provides insights into rediploidization. *Nature* **533**, 200–205, https://doi.org/10.1038/nature17164, http://www.nature.com/nature/journal/v533/n7602/abs/nature17164.html#supplementary-information (2016)
41. NCBI. (National Library of Medicine (US), National Center for Biotechnology Information, Bethesda (MD), 2002).
42. Beavis, W. D. In *Molecular dissection of complex traits* (ed A. H. Paterson) 145-162 (CRC Press, New York, 1998).
43. Koch, E., Ristroph, M. & Kirkpatrick, M. Long Range Linkage Disequilibrium across the Human Genome. *PLoS One* **8**, e80754, https://doi.org/10.1371/journal.pone.0080754 (2013).
44. Pruitt, K. D., Tatusova, T., Klimke, W. & Maglott, D. R. NCBI Reference Sequences: current status, policy and new initiatives. *Nucleic Acids Research* **37**, D32–D36, https://doi.org/10.1093/nar/gkn721 (2009).
45. Castley, A. S. L. & Martinez, O. P. In *Immunogenetics: Methods and Applications in Clinical Practice* (eds T. Frank Christiansen & D. Brian Tait) 159-171 (Humana Press, 2012).
46. Tsai, H.-Y. *et al*. Genomic prediction of host resistance to sea lice in farmed Atlantic salmon populations. *Genetics Selection Evolution* **48**, 47, https://doi.org/10.1186/s12711-016-0226-9 (2016).
47. Bangera, R., Correa, K., Lhorente, J. P., Figueroa, R. & Yáñez, J. M. Genomic predictions can accelerate selection for resistance against Piscirickettsia salmonis in Atlantic salmon (*Salmo salar*). *BMC Genomics* **18**, 121, https://doi.org/10.1186/s12864-017-3487-y (2017).
48. Wolc, A. *et al*. Breeding value prediction for production traits in layer chickens using pedigree or genomic relationships in a reduced animal model. *Genetics Selection Evolution* **43**, 5, https://doi.org/10.1186/1297-9686-43-5 (2011).
49. Daetwyler, H. D., Swan, A. A., van der Werf, J. H. J. & Hayes, B. J. Accuracy of pedigree and genomic predictions of carcass and novel meat quality traits in multi-breed sheep data assessed by cross-validation. *Genetics, Selection, Evolution: GSE* **44**, 33–33, https://doi.org/10.1186/1297-9686-44-33 (2012).
50. Chen, L., Schenkel, F., Vinsky, M., Crews, D. H. & Li, C. Accuracy of predicting genomic breeding values for residual feed intake in Angus and Charolais beef cattle1. *Journal of Animal Science.* **91**, 4669–4678, https://doi.org/10.2527/jas.2013-5715 (2013).
51. Bangera, R., Correa, K., Lhorente, J. P., Figueroa, R. & Yáñez, J. M. Genomic predictions can accelerate selection for resistance against Piscirickettsia salmonis in Atlantic salmon (*Salmo salar*). *BMC Genomics* **18**, 121, https://doi.org/10.1186/s12864-017-3487-y (2017).

52. Vallejo, R. L. *et al*. Genomic selection models double the accuracy of predicted breeding values for bacterial cold water disease resistance compared to a traditional pedigree-based model in rainbow trout aquaculture. *Genetics Selection Evolution* **49**, 17, https://doi.org/10.1186/s12711-017-0293-6 (2017).
53. Lee, J. *et al*. Comparison of alternative approaches to single-trait genomic prediction using genotyped and non-genotyped Hanwoo beef cattle. *Genetics, Selection, Evolution: GSE* **49**, 2, https://doi.org/10.1186/s12711-016-0279-9 (2017).
54. Bartölke, R., Heinisch, Jürgen, J., Wieczorek, H. & Vitavska, O. Proton-associated sucrose transport of mammalian solute carrier family 45: an analysis in Saccharomyces cerevisiae. *Biochemical Journal* **464**, 193–201, https://doi.org/10.1042/bj20140572 (2014).
55. Martínez, R. *et al*. Analysis of the expression of Kv10.1 potassium channel in patients with brain metastases and glioblastoma multiforme: impact on survival. *BMC Cancer* **15**, 1–9, https://doi.org/10.1186/s12885-015-1848-y (2015).

## Acknowledgements

The authors thank scorers (A, B and C) for the phenotyping of the fish. We also thank services from university of Oslo in the form of AEBL computer cluster. This project was mainly funded by the Norwegian Research Council under the Grant no. 235783, with additional support from Marine Harvest AS. We also thank reviewers of this manuscript for their suggestions and comments which certainly helped us to improve the manuscript.

## Author contributions

M.L.A. and S.A.B. had equal involvement with contribution in data analyses and drafting of manuscript. B.G. and M.L.H. assisted with project/study design and coordination. A.N. assisted in arranging population material, collection of data and coordination at the field recording facility. M.L.A. and S.A.B. wrote the manuscript and all other authors gave suggestions and comments for the improvement of paper. All authors read and approved the final manuscript. Overall coordination of the project was done by B.G.

## Competing interests

The authors declare no competing interests.

## Additional information

**Supplementary information** is available for this paper at https://doi.org/10.1038/s41598-020-63423-8.

**Correspondence** and requests for materials should be addressed to M.L.A.

**Reprints and permissions information** is available at www.nature.com/reprints.

**Publisher's note** Springer Nature remains neutral with regard to jurisdictional claims in published maps and institutional affiliations.